

ECE 901 HW 7: Sparse De-noising Lower Bound

Gautam Dasarathy

1 Minimax Lower bound via hypothesis testing

In this homework, you will prove a general minimax lower bound and apply it to sparse signal estimation. There are 8 problems you need to solve, indicated in **red**, below. Recall the setup we had in class:

1. \mathcal{F} is a class of models.
2. To each $f \in \mathcal{F}$, we have an associated probability distribution P_f on $\mathcal{X} \times \mathcal{Y}$ and the data $Z = (X_1, Y_1, \dots, X_n, Y_n) \in \mathcal{Z}$ is drawn *iid* from one of these distributions, i.e., $(X_i, Y_i) \stackrel{iid}{\sim} P_f, i = 1, 2, \dots, n$.
3. $d(\cdot, \cdot)$ is a semi-metric on \mathcal{F} .

What follows next is a minimax lower bounding technique based on a multiple-hypothesis test. Fill in the missing steps in the proof of the following theorem.

Theorem 1. *Under the above setup, suppose that for $M \geq 2$ there exists a set of models $\{f_1, \dots, f_M\} \subset \mathcal{F}$ such that $d(f_i, f_j) \geq 2\epsilon_n, i \neq j$, and that $\sum_{i,j=1}^M \frac{1}{M^2} D(P_{f_i} \| P_{f_j}) \leq \frac{1}{2} \log M - 1$, then the following minimax lower bound holds for the model class $\{(f, P_f) : f \in \mathcal{F}\}$.*

$$\inf_{\hat{f}_n} \sup_{f \in \mathcal{F}} \mathbb{E}_{P_f} [d(\hat{f}_n, f)] \geq \frac{\epsilon_n}{2} \quad (1)$$

Proof. We showed in class that under the hypotheses given we have that

$$\begin{aligned} \inf_{\hat{f}_n} \sup_{f \in \mathcal{F}} \mathbb{E}_{P_f} [d(\hat{f}_n, f)] &\geq \epsilon_n \inf_{\hat{f}_n} \sup_{f \in \mathcal{F}} P_f(d(\hat{f}_n, f) \geq \epsilon_n) \geq \epsilon_n \inf_{\hat{f}_n} \max_{j \in \{1, \dots, M\}} P_{f_j}(d(\hat{f}_n, f_j) \geq \epsilon_n) \\ &\geq \epsilon_n \inf_{\hat{h}_n} \max_{j \in \{0, \dots, M\}} P_{f_j}(\hat{h}_n(Z) \neq j) \geq \epsilon_n \inf_{\hat{h}_n} \frac{1}{M} \sum_{j \in \{1, \dots, M\}} P_{f_j}(\hat{h}_n(Z) \neq j), \end{aligned}$$

where $\hat{h}_n : \mathcal{Z} \rightarrow \{1, \dots, M\}$ is any hypothesis test. (Note that in class, our subset was indexed as $\{f_0, \dots, f_M\}$; in this homework, we will start the indexing from 1 instead.)

The rest of the proof will focus on obtaining the following lower bound on the average probability of error

$$\bar{P}_{e,M} \triangleq \inf_{\hat{h}_n} \frac{1}{M} \sum_{j \in \{1, \dots, M\}} P_{f_j}(\hat{h}_n(Z) \neq j) \geq \frac{1}{2}.$$

1/2 is indeed an arbitrary constant but it will suffice for our purposes and it will keep the presentation tidy.

1. Let $J \sim \text{unf}(\{1, \dots, M\})$ be a random variable drawn independently of everything else. Show that if the conditional distribution of Z given $J = j$ is P_{f_j} , then $\bar{P}_{e,M} = \mathbb{P}(\hat{h}_n(Z) \neq J)$.

Background on Information-Theoretic Concepts. Before we proceed, we need the following definitions and properties:

1. For a random variable $X \sim P_X$ the entropy is defined as $H(X) \triangleq \mathbb{E}_{P_X}[-\log P_X]$. This quantity measures the “amount of randomness” in the random variable X . If X takes values in a finite set \mathcal{X} , we have that $0 \leq H(X) \leq \log |\mathcal{X}|$; $H(X) = 0$ if X is deterministic and $H(X) = \log |\mathcal{X}|$ if $X \sim \text{uniform}(\mathcal{X})$, i.e., a uniformly distributed random variable has the most “amount of randomness”.

If $X \sim \text{Bernoulli}(p)$, then we write $h_b(p)$ to denote the entropy $H(X)$, i.e., $h_b(p) \triangleq -p \log p - (1-p) \log(1-p)$.

2. Given a pair of random variables $X, Y \sim P_X P_{Y|X} = P_{XY}$, notice that the entropy satisfies the following:

$$H(X, Y) \triangleq \mathbb{E}_{P_{XY}}[-\log P_{XY}] = \mathbb{E}_{P_X}[-\log P_X] + \mathbb{E}_{P_{XY}}[-\log P_{Y|X}].$$

The term $\mathbb{E}_{P_{XY}}[-\log P_{Y|X}]$ is denoted as $H(Y | X)$ and is called the *conditional entropy of Y given X* . It measures the average amount of randomness “left-over” in Y given the value that X takes. This property can be generalized to show that given n random variables X_1, \dots, X_n , we have the following *chain rule* of entropy

$$H(X_1, \dots, X_n) = H(X_1) + \sum_{i=2}^n H(X_i | X_1, \dots, X_{i-1}).$$

3. Given two random variables $X, Y \sim P_{XY}$, the KL divergence between their joint distribution P_{XY} and the product of their marginals $P_X \times P_Y$ is called the *mutual information between X and Y* and is denoted as $I(X; Y)$. That is

$$I(X; Y) \triangleq D(P_{XY} \| P_X P_Y).$$

(Recall that the KL divergence of P_2 from P_1 is defined as $D(P_1 \| P_2) = \mathbb{E}_{P_1}[\log(P_1/P_2)]$).

From the above definitions, it can be seen that $I(X; Y) = H(X) - H(X | Y) = H(Y) - H(Y | X)$. Also, since the KL divergence is always positive, we have that $H(X) - H(X | Y) = I(X; Y) \geq 0$ which implies that $H(X) \geq H(X | Y)$, i.e., conditioning can never increase the amount of randomness.

We will proceed by defining the following “error event” random variable

$$E = \begin{cases} 1 & \text{if } \hat{h}_n(Z) \neq J \\ 0 & \text{o.w.} \end{cases} \quad (2)$$

Let us expand the conditional entropy $H(E, J | \hat{h}_n(Z))$ in two different ways as follows:

$$H(E, J | \hat{h}_n(Z)) \stackrel{(a)}{=} H(J | \hat{h}_n(Z)) + H(E | J, \hat{h}_n(Z)) \stackrel{(b)}{=} H(J | \hat{h}_n(Z)), \quad (3)$$

2. Explain why equalities (a) and (b) above are true.

On the other hand, we have

$$H(E, J | \hat{h}_n(Z)) \stackrel{(a)}{=} H(E | \hat{h}_n(Z)) + H(J | E, \hat{h}_n(Z)) \stackrel{(b)}{\leq} h_b(\bar{P}_{e,M}) + H(J | E, \hat{h}_n(Z)),$$

3. Prove equality (a) and inequality (b) above.

$$\begin{aligned} H(J | E, \hat{h}_n(Z)) &= (1 - \bar{P}_{e,M})H(J | E = 0, \hat{h}_n(Z)) + \bar{P}_{e,M}H(J | E = 1, \hat{h}_n(Z)) \\ &\leq \bar{P}_{e,M} \log M, \end{aligned} \quad (4)$$

where we have used the fact that $H(J | E = 0, \hat{h}_n(Z)) = 0$ since given that no error occurs ($E = 0$), the value of $\hat{h}_n(Z)$ completely determines the value of J . Also, given that an error does occur ($E = 1$) and the value $\hat{h}_n(Z)$, $H(J | E = 1, \hat{h}_n(Z)) = \log(M - 1)$ since J is uniformly distributed on the set $\{1, \dots, M\} \setminus \{\hat{h}_n(Z)\}$

Combining (3) and (4), we have

$$1 + \bar{P}_{e,M} \log M \geq H(J | \hat{h}_n(Z)) \geq H(J | Z),$$

where the last inequality follows since $\hat{h}_n(Z)$ is a deterministic function of Z , and thus conditioning on Z provides at least as much information as $\hat{h}_n(Z)$. This is the so-called *data processing inequality*. The entropy (amount of randomness) in J can only be reduced if we are given Z instead of $\hat{h}_n(Z)$, since it might be the case that a deterministic function $\hat{h}_n(Z)$ does not have as much “information” about J as Z does.

Now, using the fact that $I(J; Z) = H(J) - H(J | Z) = \log M - H(J | Z)$, we have the following inequality (which is a version of Fano’s inequality)

$$\bar{P}_{e,M} \geq 1 - \frac{I(J; Z) + 1}{\log M}. \quad (5)$$

Finally, we observe that

$$\begin{aligned}
I(J; Z) &\stackrel{(a)}{=} D(P_{JZ} \| P_J P_Z) \\
&\stackrel{(b)}{=} \frac{1}{M} \sum_{i=1}^M D \left(P_{f_i} \left\| \sum_{j=1}^M P_{f_j} \right. \right) \\
&\stackrel{(c)}{\leq} \frac{1}{M^2} \sum_{i,j=1}^M D(P_{f_i} \| P_{f_j})
\end{aligned}$$

where (a) and (b) follow from the definition, and (c) follows from the fact that the KL divergence is convex.

4. Show that this last inequality, when used in (5), concludes the proof. □

2 Minimax Lower Bound for Sparse Denoising

Consider the problem of estimating a vector $f \in \mathbb{R}^n$ from noisy measurements $z \in \mathbb{R}^n$ of the form

$$z = f + \epsilon, \tag{6}$$

where $\epsilon \sim \mathcal{N}(0, \mathbf{I})$ and $f \in \mathcal{F}_k \triangleq \{x \in \mathbb{R}^n : \|x\|_0 \leq k\}$, i.e., f has at most k non-zero entries.

We saw in class that, using soft-thresholding, it is possible to achieve a mean-squared error upper bound of

$$(2 \log n + 1) \left[1 + \sum_{i=1}^n \min \{ |f(i)|^2, 1 \} \right] = (2 \log n + 1) [1 + k],$$

where $f(i)$ is the i -th coordinate of the vector f . Notice that the quantity on the right is no greater than $4k \log n$ if $n, k \geq 3$.

Your mission, if you choose to accept it, is to prove the following minimax lower bound: There exists a constant $C > 0$ such that

$$\mathcal{E}^* \triangleq \inf_{\hat{f}} \sup_{x \in \mathcal{F}_k} \mathbb{E} \left[\|\hat{f}(z) - f\|_2^2 \right] \geq C \max \left\{ k \log \left(\frac{n}{k} \right), k \right\}, \tag{7}$$

where \hat{f} is any estimator. This shows that the $\log n$ factor is an unavoidable consequence of having to identify the unknown pattern of sparsity. You can use the following outline. First, prove the following simple lower bound.

5. Show that k is a lower bound on the mean squared error. (Hint: Suppose we are given the support of f .)

The lower bound of k is loose when $k \ll n$, and a tighter lower bound can be obtained using multiple hypothesis testing (Theorem 1). The idea behind using the above theorem is to identify a set of M examples that demonstrate the hardness of the problem. This will be your first task:

Assume that n and k are fixed, and for simplicity, assume that k is even and that $k < n/2$. We will consider k -sparse vectors with non-zero elements equal to $a_n > 0$ in magnitude (will choose a_n later to maximize the lower bound). Show there exists a set of such vectors $\{f_1, \dots, f_M\} \subset \mathcal{F}_k$, with $M = \left(\frac{n}{k}\right)^{k/4}$, satisfying

$$\|f_i - f_j\|_2^2 \geq a_n/2, \quad \forall i \neq j. \quad (8)$$

While it is possible to show an explicit construction of this set, you can instead use the following outline which employs the probabilistic method.

6. Consider a set $\mathcal{U} = \{f \in \{-\sqrt{\frac{a_n}{k}}, 0, +\sqrt{\frac{a_n}{k}}\}^n : \|f\|_0 = k\}$. Show that the cardinality of this set is $|\mathcal{U}| = \binom{n}{k} 2^k$. Also, show that for a fixed $f \in \mathcal{U}$, the following holds

$$|\{f' \in \mathcal{U} : \|f - f'\|_0 \leq k/2\}| \leq \binom{n}{k/2} 3^{k/2}.$$

7. Use (i) to argue that if one picks a size $(n/k)^{k/4}$ set of elements at random from \mathcal{U} , then with probability greater than $1/2$, this random set has the property (8). For this you might find the following combinatorial inequality (which holds when $n - k > (0.5)k$) useful:

$$\frac{\binom{n}{k}}{\binom{n}{k/2}} \geq \left(\frac{n}{k} - \frac{1}{2}\right)^{k/2}.$$

8. Argue that this shows that there exists at least one set of vectors $\{f_1, \dots, f_M\}$ such that $M + 1 = \left(\frac{n}{k}\right)^{k/4}$ which satisfies property.

9. Conclude the proof of the theorem using Theorem 1. (Hint: Compare the results of (a) and (b) with the hypotheses of Theorem 1; this should tell you the right value of a_n .)