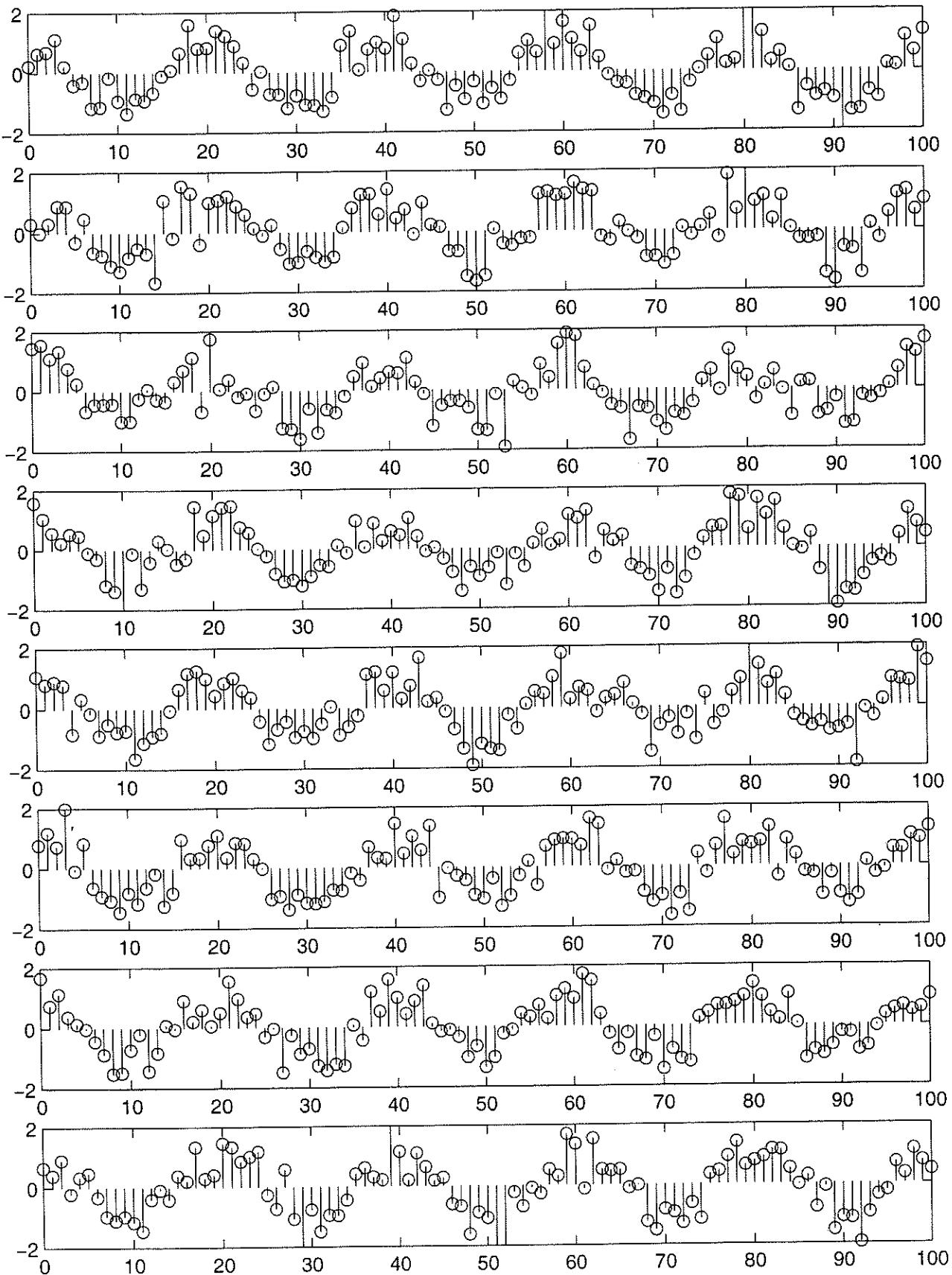


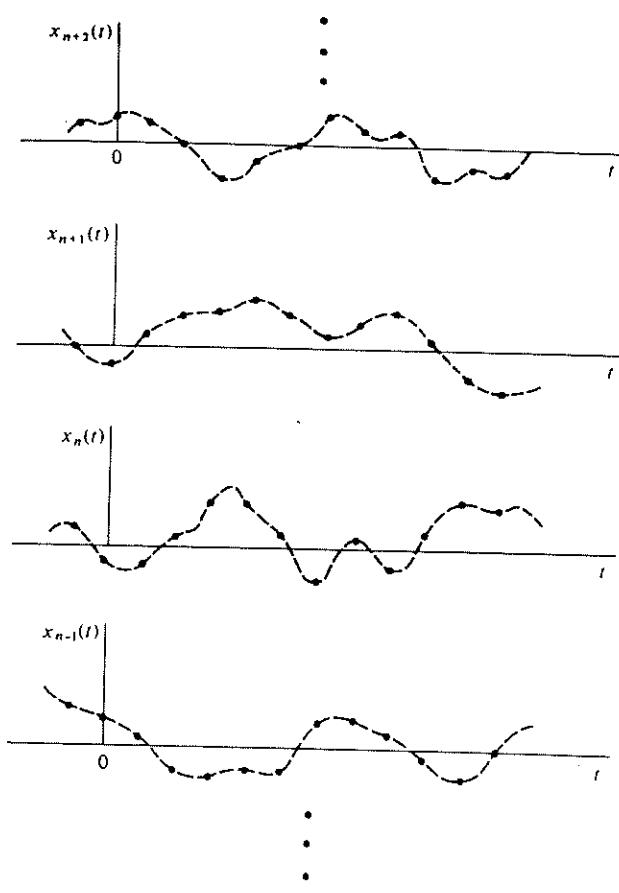
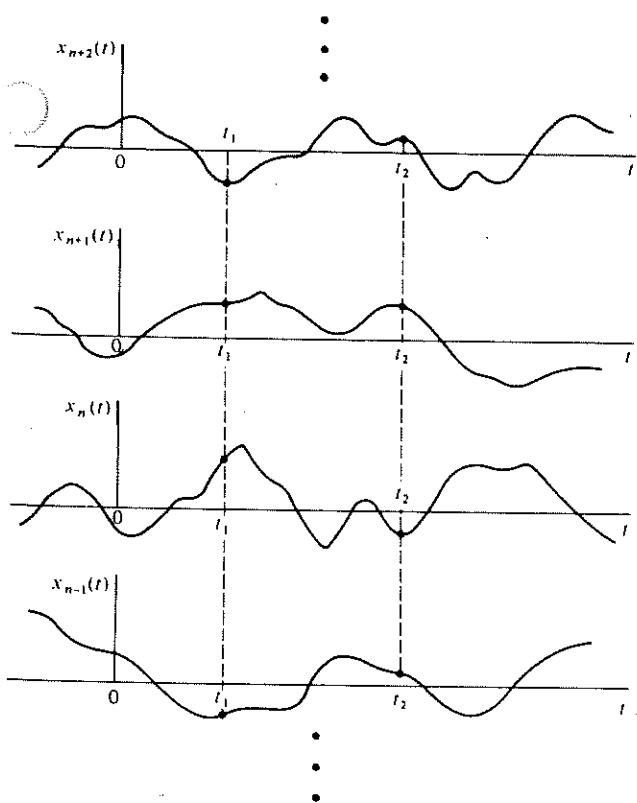
# Discrete-Time Random Signals

# Classic Example: Sinusoid + Noise

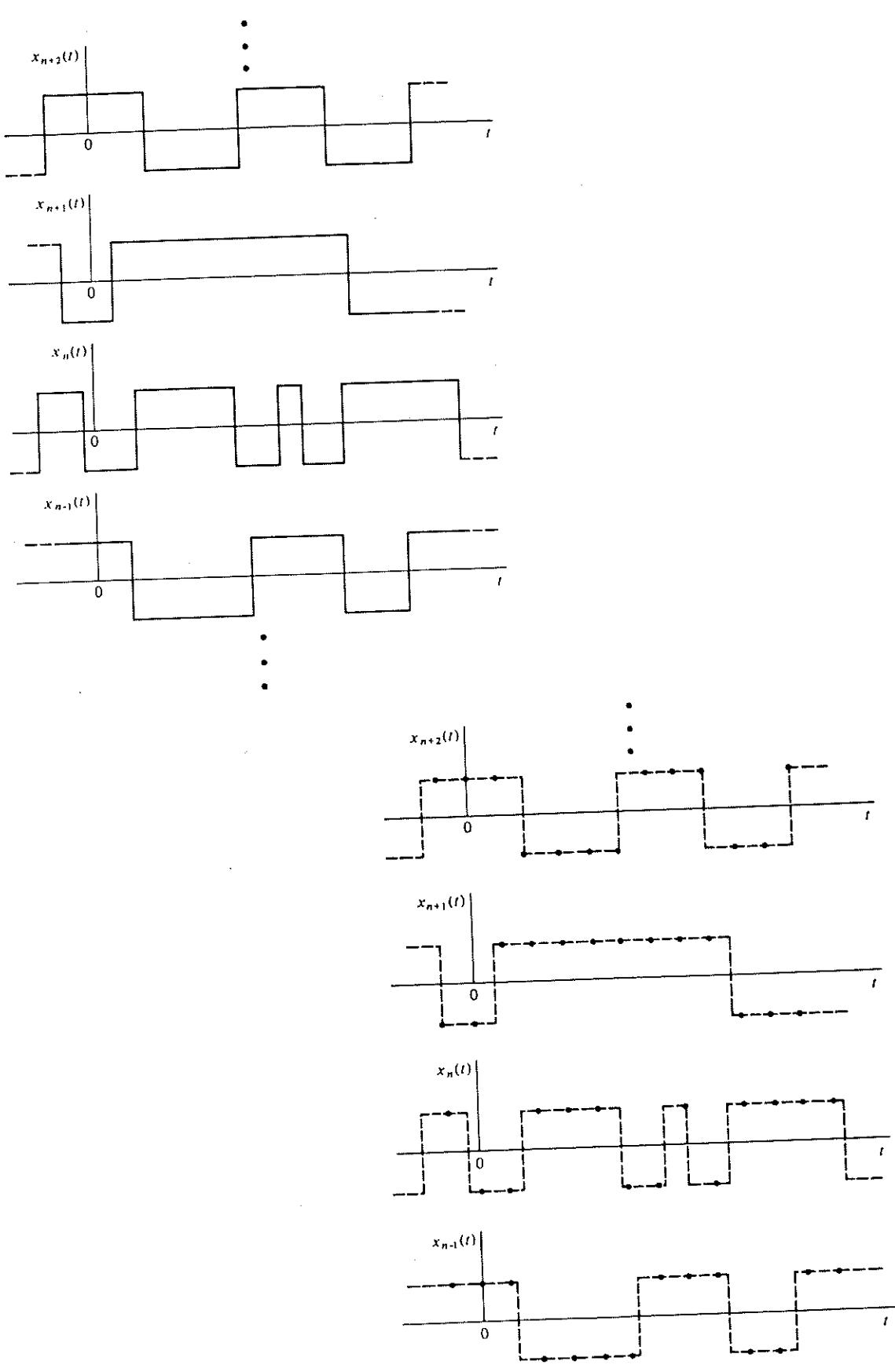


Definition: A random process is a family or ensemble of signals corresponding to every possible outcome of a certain signal measurement or experiment. Each signal in the ensemble is called a "realization" of the process.

Ex.



# Ex. Random Binary Process

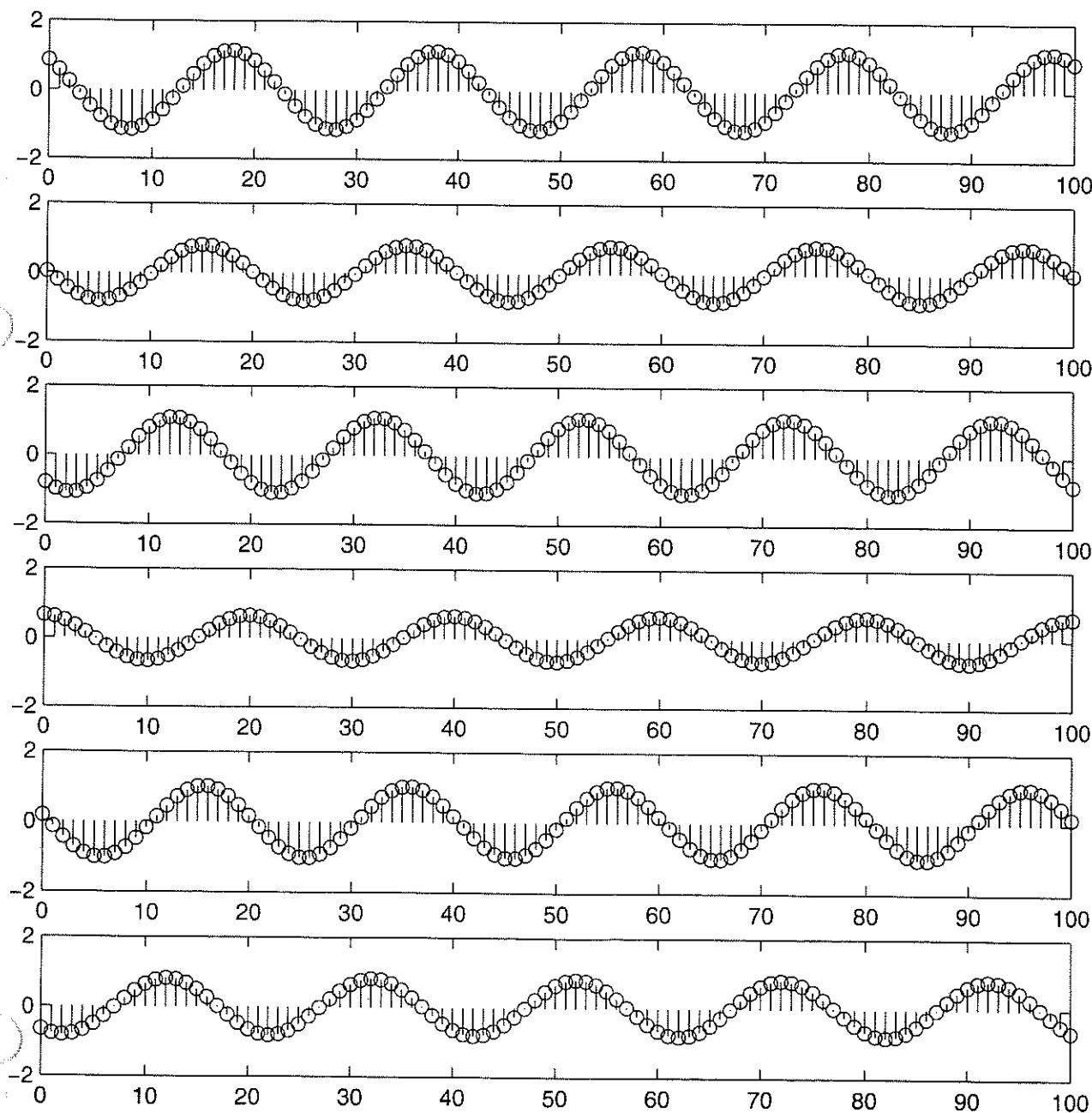


## Ex. Random Sinusoidal Process

$$x[n] = A \cos(\omega_0 n + \phi)$$

$A, \omega_0, \phi$  may all be random variables

random amplitude  
& phase



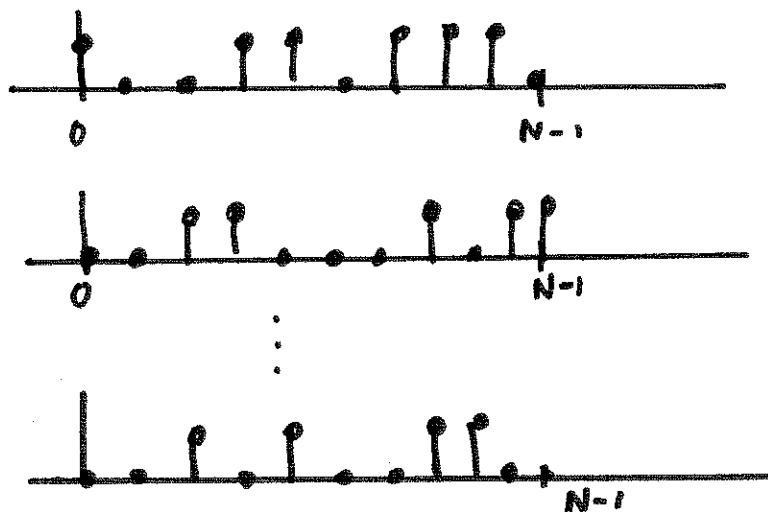
Definition: The mean of a random process is the average of all realizations of the process.

Ex. N-pt binary random sequence

$$x[n] = \begin{cases} 1, & \text{with prob } p \\ 0, & \text{with prob } (1-p) \end{cases} \quad \begin{matrix} n=0, \dots, N-1 \\ \text{independent} \end{matrix}$$

$$\Pr(x[n]=1) = p \quad \Pr(x[n]=0) = 1-p$$

Realizations:



What is  $\text{mean}(x[n])$ ?

The mean is also known as the expectation, and is denoted by

$$m_x[n] = E[x[n]]$$

↑  
expectation operator

$\Rightarrow$  take average of all  
possible realizations  
of  $x[n]$

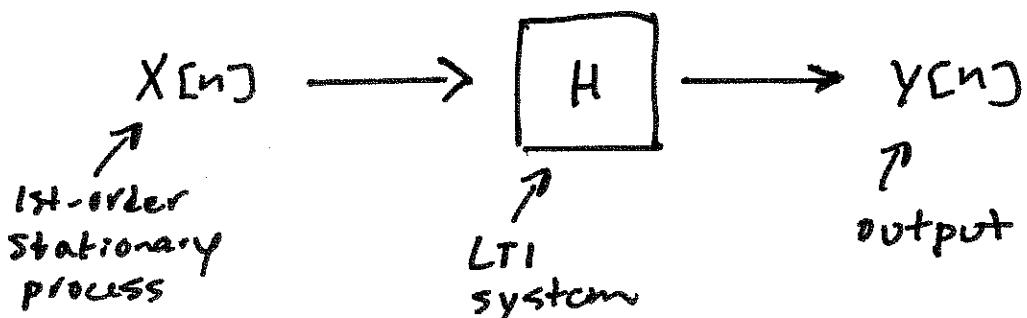
### First-order Stationary Processes

A process is 1st-order stationary  
if

$$m_x[n] = m_x \text{ (a constant independent of } n)$$

Ex. N-pt binary process revisited

## First-order Stationarity and LTI Systems



LTI system is deterministic, but input  
is random.

What about output?

$$m_y[n] = E \left[ \sum_{k=-\infty}^{\infty} h[k] x[n-k] \right]$$

=

Definition: The autocorrelation function

of a random process is the average product of a signal realization with a time-shifted version of itself.

$$R_{xx}[n, n+m] = E[x[n]x[n+m]]$$

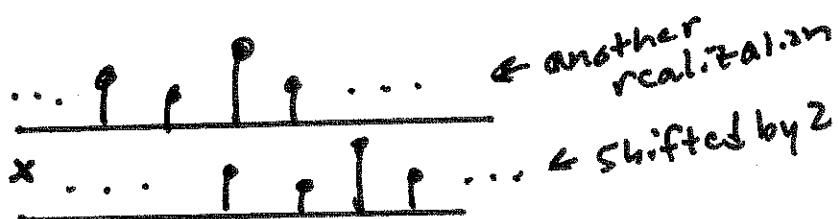
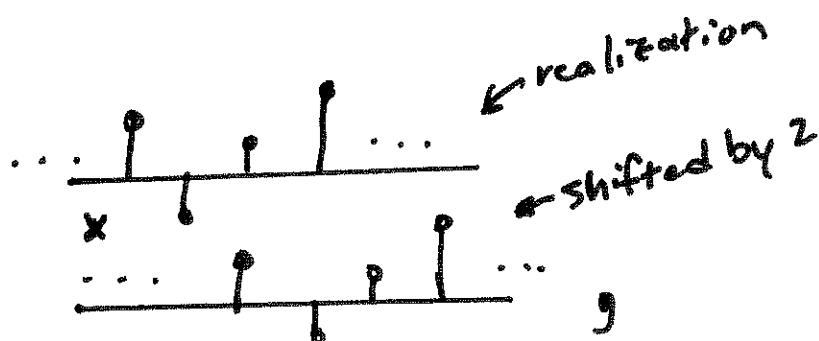
↑  
Average  
over all possible realizations

Autocorrelation  
function

Ex.

$$R_{xx}[n, n+2] =$$

Average {



⋮      }

### Ex. Random Binary Process

Assume  $\Pr(x[n] = 1) = \Pr(x[n] = -1) = \frac{1}{2}$

$$\Rightarrow m_x[n] =$$

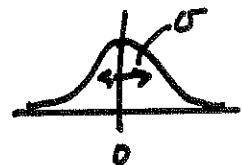
$$R_{xx}[n, n+m] = E[x[n]x[n+m]]$$

## Ex. Gaussian White Noise (GWN)

⇒ "white"  $\Rightarrow x[n]$  are independent  
and zero-mean

$$\Pr[a \leq x[n] \leq b] = \underbrace{\int_a^b \frac{1}{\sqrt{2\pi\sigma^2}} e^{-x^2/2\sigma^2} dx]_{}^{}}$$

Gaussian density function  
 $\sigma^2$  = variance



$$m_x = E[x[n]] = \int_{-\infty}^{\infty} x \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-x^2/2\sigma^2} dx = 0$$

$$R_{xx}[n, n+m] = E[x[n] x[n+m]]$$

=

## Second-order Stationary Processes

A random process is 2nd-order stationary if

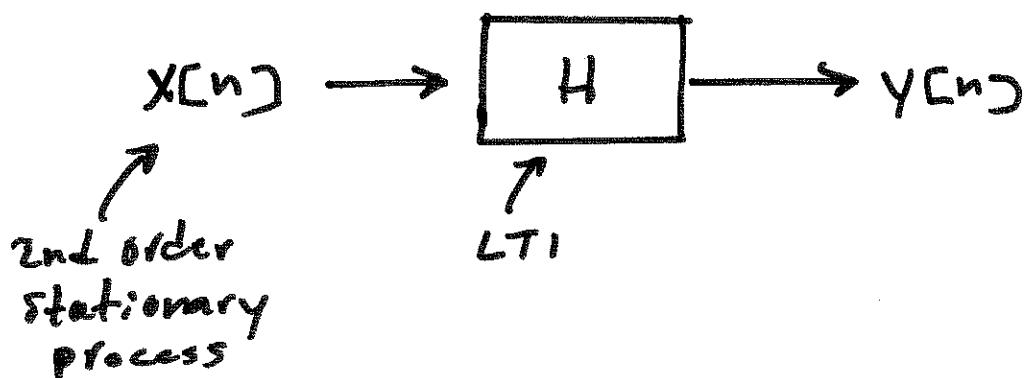
$$R_{xx}[n, n+m] \equiv R_{xx}[m]$$

That is, the autocorrelation function only depends on  $m$ , the shift, and not on  $n$ .

Ex. Binary Process

Ex. GWN

# Stationary Inputs $\notin$ LTI Systems



$$\begin{aligned} m_y(n) &= E \left[ \sum_{k=-\infty}^{\infty} h[k] x[n-k] \right] \\ &= \sum_{k=-\infty}^{\infty} h[k] E[x[n-k]] = m_x \sum_{k=-\infty}^{\infty} h[k] \end{aligned}$$

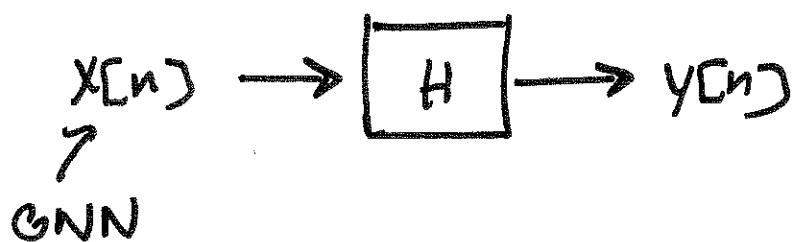
$$R_{yy}(n, n+m) = E \left[ \sum_{k=-\infty}^{\infty} h[k] x[n-k] \cdot \sum_{r=-\infty}^{\infty} h[r] x[n+m-r] \right]$$

=

$$\Rightarrow R_{yy}[m] = \sum_{l=-\infty}^{\infty} R_{xx}[m-l] \cdot c_{hh}[l]$$

where  $c_{hh}[l] = \sum_{k=-\infty}^{\infty} h[k] h[l+k]$

## Ex. GWN into an LTI System



$$m_y = m_x \sum_{k=-\infty}^{\infty} h[k] = 0$$

$$R_{yy}[m] = \sum_{l=-\infty}^{\infty} R_{xy}[m-l] \cdot \sum_{k=-\infty}^{\infty} h[k] h[l+k]$$

$\sigma^2 \delta[m-l]$

$$= \sum_{l=-\infty}^{\infty} \sigma^2 \delta[m-l] \cdot \sum_{k=-\infty}^{\infty} h[k] h[l+k]$$

$$= \sum_{k=-\infty}^{\infty} \sigma^2 h[k] h[m+k]$$

# Frequency Domain Analysis

(assume 2nd-order stationarity)

## Power Spectrum

$$S_{xx}(\omega) = \sum_{m=-\infty}^{\infty} R_{xx}[m] e^{-j\omega m} \quad \textcircled{1}$$

$$R_{xx}[m] = \frac{1}{2\pi} \int_{-\infty}^{\infty} S_{xx}(\omega) e^{j\omega m} d\omega \quad \textcircled{2}$$

① & ② are called the Wiener-Khinchin Relations

## Frequency Domain Analysis of LTI Systems

$$R_{yy}[m] = \sum_{\ell=-\infty}^{\infty} R_{xx}[m-\ell] C_{hh}[\ell]$$

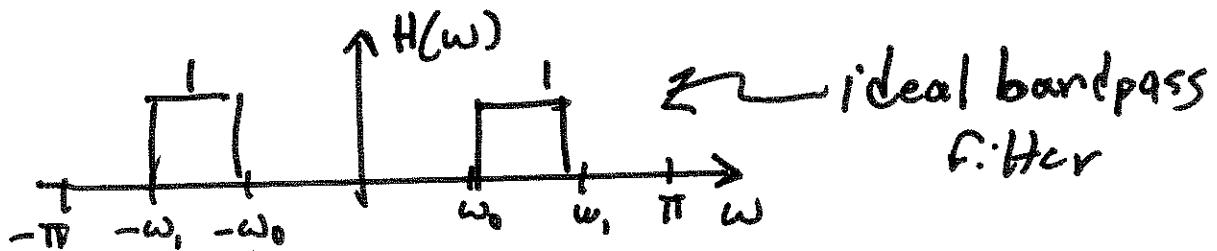
$$\Rightarrow S_{yy}(\omega) = |H(\omega)|^2 \cdot S_{xx}(\omega)$$



Power spectrum of  $X[n]$  is  
"shaped" by  $|H(\omega)|^2$

Ex.

$$H(\omega) = \begin{cases} 1, & \omega_0 \leq |\omega| \leq \omega_1 \\ 0, & \text{otherwise} \end{cases}$$



$$S_{yy}(\omega) = |H(\omega)|^2 S_{xx}(\omega)$$

In particular, the average output power is

$$\begin{aligned} E[\hat{z}_n] &= R_{yy}[0] = \frac{1}{2\pi} \int_{-\pi}^{\pi} S_{yy}(\omega) d\omega \\ &= \frac{1}{2\pi} \int_{-\omega_1}^{-\omega_0} S_{xx}(\omega) d\omega + \frac{1}{2\pi} \int_{\omega_0}^{\omega_1} S_{xx}(\omega) d\omega \end{aligned}$$

This shows that we can interpret the area under  $S_{xx}(\omega)$  for  $\omega_0 \leq |\omega| \leq \omega_1$  as the average input power in that frequency band. Thus,  $S_{xx}(\omega)$  can be viewed as a density function for power in the spectral (freq) domain.

## Applications

### Ex. Noise Removal

Suppose that we make noisy measurements of a random signal  $s[n]$  in GWN  $w[n]$ :

$$x[n] = s[n] + w[n]$$

signal:

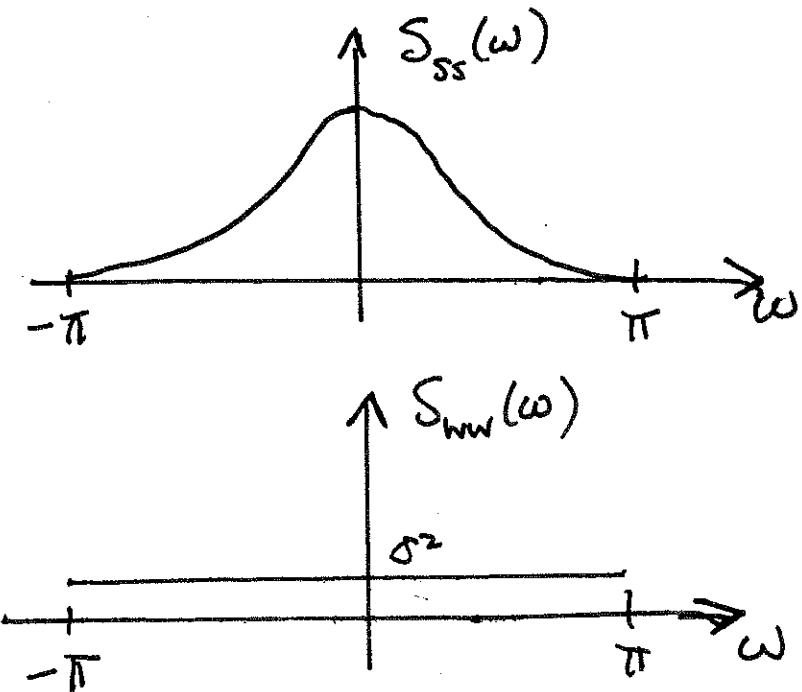
$$m_s = 0 \quad R_{ss}[m], S_{ss}(\omega)$$

arbitrary

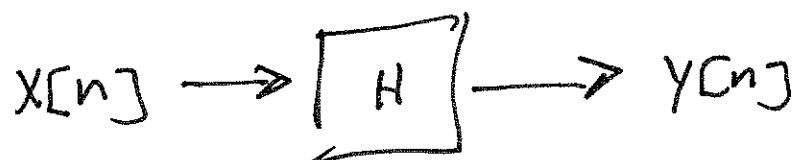
noise:

$$m_w = 0 \quad R_{ww}[m] = \sigma^2 \delta[m]$$
$$S_{ww}(\omega) = \sigma^2 \text{ const.}$$

Consider the case



Define the noise removal filter

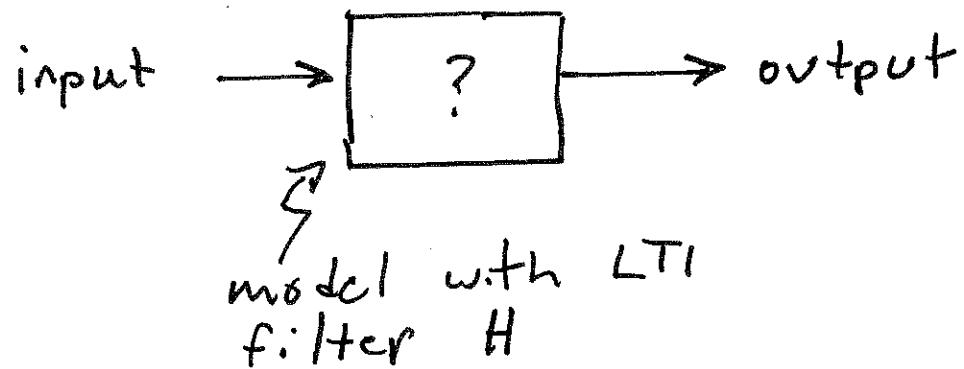


$$H(\omega) = \begin{cases} 1, & |\omega| < \omega_c \\ 0, & \omega_c \leq |\omega| \leq \pi \end{cases}$$

How to choose  $\omega_c$ ?

## Ex. System Identification

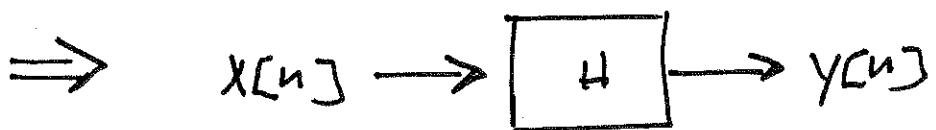
Suppose that we wish to model an unknown physical system with an LTI filter.



How can we "identify" the filter  $H$ ?

# White Noise Probing:

$$\text{input} \equiv x[n] = \text{GWN}$$



$$m_y = 0$$

$$R_{yy}[m] = \sum_{k=-\infty}^{\infty} \sigma^2 h[k] h[m+k]$$

$$S_{yy}(\omega) = \sigma^2 |H(\omega)|^2$$

close, but  
only tells us  
magnitude

# Stationary Processes

## Properties of Correlation Function:

(1)  $|R_{xx}[m]| \leq R_{xx}[0]$

(2)  $R_{xx}[-m] = R_{xx}(m)$

(3) power of process  $x[n]$  is given by  $R_{xx}[0]$

proof of (1) :  $(x[n] - x[n+m])^2 \geq 0$   
 $\Rightarrow E[(x[n] - x[n+m])^2] \geq 0$

similarly,  $(x[n] + x[n+m])^2 \geq 0$   
 $\Rightarrow E[(x[n] + x[n+m])^2] \geq 0$

proof of (2) :

$$R_{xx}[m] = E[x[n]x[n+m]] \quad \text{for any } n!$$

$$R_{xx}[-m] = E[x[n]x[n-m]]$$

take  $n = n' - m$

$$R_{xx}[m] = E[x[n'-m]x[n']] = R_{xx}[-m]$$

## Correlation

Total Correlation :



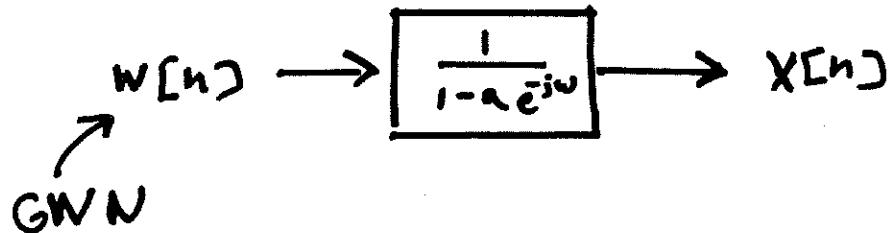
$$x[n] = A, \quad P_A(a) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{a^2}{2\sigma^2}}$$

Mean :

correlation function :

Zero Correlation :

## Strong Correlation:



$$x[n] = \alpha x[n-1] + w[n] ; 0 < \alpha < 1$$

impulse response  $h[n] =$

mean:

correlation function:

## Stationarity

Let  $x[n]$  be a random process.

Then we can talk about

$$\Pr(|x[n]| < a)$$

or  $\Pr(a < x[n] < b)$ .

We can also study the joint probability distributions for two or more samples:

$$\Pr(a < x[n] < b, c < x[n+m] < d)$$

or  $\Pr(|x[n]| \leq a, |x[n+1]| \leq b, |x[n+2]| \leq c)$

## Strict-Sense Stationarity (SSS)

Strict-sense stationarity means that these joint probability distributions are time-invariant.

ex. SSS

$$\begin{aligned} \Pr(a < x[n] < b, c < x[n+m] < d) \\ = \Pr(a < x[n+k] < b, c < x[n+k+m] < d) \\ \text{for every } k \in \mathbb{Z}. \end{aligned}$$

SSS  $\neq$  LTI Systems

$$x[n] \xrightarrow{\text{SSS}} [H] \rightarrow y[n]$$

$$y[n] = \sum_{k=-\infty}^{\infty} h[k] x[n-k]$$

$$\begin{aligned} \Pr(a \leq y[n] \leq b) &= \Pr(a \leq \sum_k h[k] x[n-k] \leq b) \\ &= \Pr(a \leq \sum_k h[k] x[n+m-k] \leq b) \quad \text{by SSS of } x[n] \\ &= \Pr(a \leq y[n+m] \leq b) \end{aligned}$$

## Wide-Sense Stationary (WSS)

$$(1) E[x[n]] = m_x \text{ const.}$$

$$(2) E[x[n]x[n+m]] = R_{xx}[m]$$

(1)  $\Rightarrow$  average value of process  
is time-invariant

(2)  $\Rightarrow$  average power of process  
and correlation between samples  
is time-invariant

WSS is weaker than SSS:

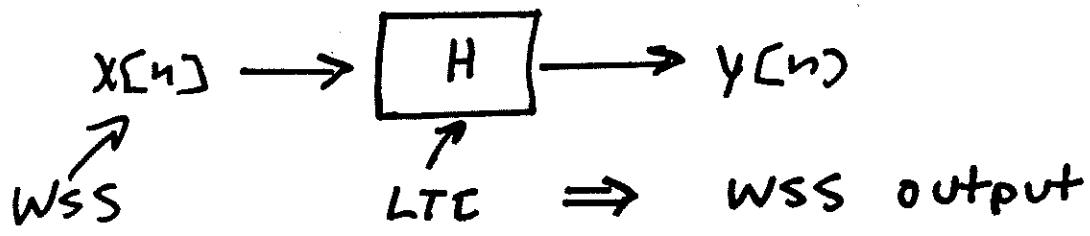
$$\text{SSS} \Rightarrow \text{WSS}$$

$\nLeftarrow$

WSS tells us that average behavior  
of single samples and products of  
samples is time-invariant.

SSS tells us that the joint probability  
distributions for arbitrary collections  
of samples are time-invariant.

## WSS & LTI Systems



In most cases and applications WSS sufficient for analysis purposes.

$$(1) E[x[n]] = m_x \Rightarrow E[y[n]] = m_y \text{ const.}$$

$$(2) E[x[n]x[n+m]] = R_{xx}[m]$$

Correlation structure  
is sufficient to describe  
frequency content of input

$$S_{xy}(\omega) = \sum_{m=-\infty}^{\infty} R_{yy}[m] e^{-j\omega m}$$

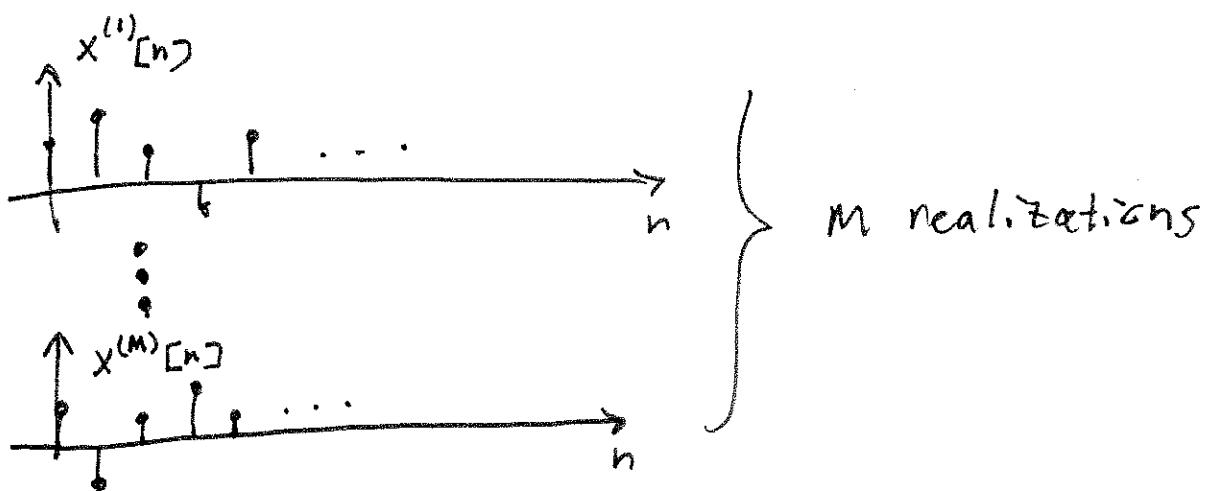
Thus, WSS is sufficient for frequency-domain analysis of random processes through LTI systems.

## Estimating Means and Autocorrelations from Data

Suppose that we make several independent observations of a random process:

$$x^{(i)}[n], \quad i = 1, \dots, M$$

realizations



We can estimate the mean function using

$$\hat{m}_x[n] = \frac{1}{M} \sum_{i=1}^M x^{(i)}[n]$$

$$\hat{m}_x[n] \rightarrow m_x[n] \text{ as } M \rightarrow \infty$$

We can estimate the autocorrelation function according to

$$\hat{R}_{xx}[n, n+m] = \frac{1}{M} \sum_{i=1}^M x^{(i)}[n] x^{(i)}[n+m]$$

If the process is stationary, then we can estimate the mean and auto correlation from a single realization using time-averaging.

$$\hat{m}_x = \frac{1}{N} \sum_{n=0}^{N-1} x[n]$$

$$\hat{R}_{xx}[m] = \frac{1}{N-m} \sum_{n=0}^{N-m-1} x[n] x[n+m]$$

$$\hat{m}_x \rightarrow m_x$$

as  $N \rightarrow \infty$

$$\hat{R}_{xx}[m] \rightarrow R_{xx}[m]$$